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# Efficacy of cognitive training on executive functions in healthy older adults: a systematic review with metaanalysis of randomized controlled trials

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#### ABSTRACT

**Objective:** Systematically review randomized controlled trials on the efficacy of cognitive training on executive functions in healthy older people.

**Measures:** The outcome measures were related to inhibitory control, working memory, and cognitive flexibility.

**Results:** Thirty-one trials were included in the systematic review and thirteen trials in the meta-analysis. In the overall analysis, the cognitive training enhanced inhibitory control when measured by the Stroop task (p < .001, d=1.64) and working memory when measured by the Corsi Block task (p=.002, d=.16). A marginal significance was found for working memory in the Digit Span task – Forward (p=.06, d=.92). However, cognitive training did not enhance inhibitory control when measured by the Go/No-Go task (p=.76, d=.59), working memory when measured by the Digit Span – Backward (p=.72, d=.95) and N-Back (p=.10, d=.26) tasks, and cognitive flexibility when measured by Trail Making – Part B (p=.08, d=.27) and Semantic Fluency (p=.49, d=.06) tasks.

**Conclusion:** Mixed evidence was found for inhibitory control and working memory; cognitive flexibility showed no evidence of improvement.

#### **ARTICLE HISTORY**

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#### **KEYWORDS**

Older adults; meta-analysis; cognitive aging; cognitive training; executive functions

The increase in human longevity, driven by improvements in living conditions, nutrition, medical technology, and cognitive development, has dramatically changed the prospects of future life, especially for older adults (Caswell & Zarulli, 2018; Maldonado Briegas et al., 2020). The increase of longevity must be seen as a modern science achievement. The United Nations estimates that the number of older adults (i.e.  $\geq$ 65 years old) worldwide will rise from 0.7 billion (9%) in 2019 to 1.5 billion (16%) in 2050 (United Nations, 2019). However, a higher proportion of older people are susceptible to illnesses since the aging process is generally described as being closely associated with the onset of many diseases (Lazarus & Harridge, 2018). The review

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proposed by Jaul and Barron (2017) summarized the main age-related diseases in older adults and highlighted mild short-term memory loss, word-finding difficulty, and slower processing speed as normal parts of aging. The neuropsychological literature indicates that healthy older adults showed worse performance than healthy younger adults on a variety of cognitive tasks that assessed: processing speed, inhibition, and visual-spatial ability (Ferguson et al., 2021; Kujawski et al., 2021; Langeard et al., 2021; Li et al., 2021). In contrast, the review proposed by Harada et al. (2013) reported that some crystallized abilities (e.g. vocabulary) show a slower decline due to brain aging and may even improve with age because of knowledge accumulated during their lives. For instance, older adults show better performance when compared to younger adults on tasks in which they use wisdom (e.g. judgment and problem-solving tasks; Dumas, 2017). Therefore, there is a great and growing effort of the neuroscience community on the nature of later life, including how to sustain cognitive health and even how to enhance it (Foster & Walker, 2021).

Cognitive functions (e.g. perception, attention, memory, problem solving, executive functions, decision making, intelligence) play a crucial role in our functioning, impacting our everyday life activities (see Demir Akça et al., 2014; Murman, 2015). Some functions—especially the executive functions—decline gradually over time as a result of the continuous aging process, that is, nonpathological and age-associated cognitive decline (Murman, 2015). The executive functions are key functions for everyday life, allowing individuals to plan ahead, focus their attention, switch between different tasks, and maintain effective levels of independent functioning (Corbo & Casagrande, 2022; Ferguson et al., 2021). The literature shows that a good performance in executive functions is associated with successful goal attainment in a variety of contexts (e.g. Pascual et al., 2019; Peng et al., 2022), independence, autonomy, and guality of life (Coppin et al., 2006; Gamage et al., 2018), as well as oriented behaviours and complex situations, such as crossing the street (Nicholls et al., 2022). Deficits in executive functions have been associated with poor outcomes, such as poor social function, impaired instrumental activities of daily living, and depression (Alexopoulos et al., 2000; Duggan et al., 2017). Studies conducted with older adults have demonstrated declines in executive cognitive functions (e.g. inhibition control, mental shifting; Peng et al., 2022) which is gradually mediated by a decrease in brain volume (in the right parietal and prefrontal cortices, for instance; Fastame et al., 2022). As a result of aging, the executive functions are one of the first cognitive functions to decline due to micro and macrostructural alterations in the brain connectivity (for a functional and structural perspective, see Fjell et al., 2017). The executive functioning is a higher-order processing activity in the brain, and it is the process by which individuals exercise conscious control over their thoughts and actions (Fan & Wang, 2023). Most related changes in executive functions are suggestive of impairment in the frontal lobes, and changes in the frontostriatal circuit (i.e. neural pathways connecting the frontal lobe with the basal ganglia) are possibly the most significant cause of impaired executive function in older adults with no dementia (Lima-Silva et al., 2012).

In terms of constituents of executive functioning, one of the most accepted theoretical frameworks proposes a core triad of functions: inhibitory control, working memory, and cognitive flexibility (see Diamond, 2013; Lehto et al., 2003; Miyake et al., 2000). Inhibitory control is the cognitive ability to suppress or countermand a thought, action, or feeling (Spechler et al., 2016). It allows an individual to inhibit their impulses and natural, habitual, or dominant behavioural responses to stimuli in order to select more appropriate behaviours consistent with one's goals (Li et al., 2022). Inhibitory control can be measured using classical paradigms of experimental psychology, for example, the Stroop task, Go/No-Go task, and the Stop-Signal task (Kang et al., 2021). Working memory is the cognitive ability that allows an individual to hold a small amount of information that can be held in mind and applied in the execution of cognitive tasks (Cowan, 2014). It is essential to all advanced thinking to learn facts or skills (Bergman Nutley & Södergvist, 2017). Experimentally, working memory can be measured using classical paradigms, for example, the Digit Span task, Letter/ Number Sequencing task, and the Corsi Block task (Shelton et al., 2009). Cognitive flexibility is the ability that allows an individual to efficiently adjust one's behaviour according to a changing environment (Dajani & Uddin, 2015). It enables individuals to integrate external evidence into previous expectancies (Romero-Ferreiro et al., 2022). Experimentally, cognitive flexibility can be measured using classical paradigms, for example, the Trail-Making Task—Part B, Wisconsin Card Sorting Task, and Fluency tasks (Takeda & Fukuzaki, 2021).

The human brain is inherently plastic and is continually adapting to its environment. Thus, despite the inevitable aging process, engagement in cognitive activities (e.g. learning a new language, maintaining social connections, and engaging in challenging cognitive tasks) can potentially mitigate cognitive declines (Stieger & Lachman, 2021). Importantly, cognitive training of executive functions seems to increase functional and neural plasticity, even in older age (Nguyen et al., 2019). Cognitive training is an intervention centred on the cognitive performance that uses a set of standardized behavioural task protocols that tackle cognitive functions (Golino & Flores-Mendoza, 2016), and that may be associated with other interventions (e.g. physical exercise; Anguera et al., 2022). These 'trainable' functions range from lower level processes—for example, perception: biological motion—to higher order processes—for example, executive functions: working memory (see Legault & Faubert, 2012; Weng et al., 2019). The efficacy is usually assessed through cognitive evaluation (e.g. neuropsychological testing) for one or several cognitive domains before and after the intervention.

Considering that older adults have a high risk of serious cognitive diseases, identification of strategies and possible interventions for preventing cognitive decline is necessary (Giuli et al., 2016). In recent times, several devices and platforms have started to play a significant role in cognitive training since such training can potentially be undertaken at any time and accessed from anywhere (Klimova, 2016). Rapid advances in computing technology had evolved exponentially over time due to a fusion of technologies that are blurring the lines between physical, digital, and biological spheres (Park, 2016). As a result, this has enabled researchers and clinical professionals to conduct accessible and fine-tuned cognitive training using virtual reality, interactive video game playing, mobile setup, and other cutting-edge technologies (Ge et al., 2018).

Over the last years, many studies showed positive results of cognitive training on cognitive functioning in older adults. A review proposed by Sanjuán et al. (2020) endorsed the effectiveness of cognitive interventions in older adults. However, the authors highlighted aspects that must be met by proper experimental protocols for

cognitive training (e.g. session length, total number of sessions, measures of daily functioning) in order to make the intervention more effective. Additionally, with the growing number of publications related to cognitive training applied in clinical populations, there has been an increase in the number of systematic reviews with and without meta-analysis. A meta-analysis proposed by Yun and Ryu (2022) demonstrated cognitive training was the most effective intervention in healthy older adults in comparison to cognitive stimulation (e.g. reality orientation) and cognitive rehabilitation (e.g. activities to improve the performance of daily activities). Nevertheless, systematic reviews show conflicting results (Makin, 2016; Traut et al., 2021). Some reviews fond clear benefits to a trained ability—for example, executive function in older adults with cognitive impairment (see Abd-alrazaq et al., 2022), while other reviews yield little to no evidence of benefit from cognitive training (see Sala et al., 2019). In addition, the reviews usually assess the efficacy of cognitive training on outcomes related to global cognition and neglect cognitive subdomains.

Therefore, and because cognitive training deals extensively with several areas (e.g. basic science, health, public policies, industry, and marketing), systematic reviews must be conducted periodically to present the state-of-the-art of the field and show its improvements in terms of methodological control. Four previous studies have performed systematic reviews with meta-analysis to address the effect of cognitive training on executive function in healthy older adults (Chiu et al., 2017; Lampit et al., 2014; Nguyen et al., 2019; Wollesen et al., 2020). Considering the high heterogeneity of the trials in terms of intervention implemented, type of controls, cognitive processes assessed, and response variables, in addition to the increase in published papers, professionals and decision makers are seeking reviews with detailed data and higher degrees of specificity. The meta-analyses conducted by Lampit et al. (2014) and Chiu et al. (2017) investigated outcomes related to different cognitive functions (memory, attention, processing speed, executive function, and visual-spatial skills) in healthy older adults. Nguyen et al. (2019) segmented the meta-analysis by executive functions subdomains (inhibitory control, working memory, and cognitive flexibility), and Wollesen et al. (2020) split the meta-analysis by type of control (active or passive) in the same population. We went further and ran task-specific meta-analyses and provided meta-analytic data of the neuropsychological tests/paradigms most used in cognitive training protocols for older people.

# Method

The current study was registered in PROSPERO (ID: CRD42021237057) and conducted in accordance with recommendations outlined by the PRISMA group guidelines (Page et al., 2021; see Supplementary Material—Appendix A for PRISMA-checklist).

# **Eligibility criteria**

We targeted the research question using the PICOS framework: Population—cognitively healthy older adults; Intervention—cognitive training to enhance or maintain executive functioning; Comparison—active or passive control groups; Outcome -neuropsychological measures of executive functions; Study design—randomized controlled trials.

# Type of studies

We first identified and then collected peer-reviewed scientific papers of trials from online electronic databases written in any language that investigated the effect of cognitive training on executive functions outcomes in cognitively healthy older adults.

# Type of participants

The total experimental sample of each trial had to comprise individuals aged 59 years and older with normal cognitive functioning, and that have not been diagnosed with mild cognitive impairment or any form of dementia (experimental and control groups). The eligibility was confirmed by examining the baseline characteristics of the sample and the trial inclusion criteria.

# Type of intervention

The intervention consisted of cognitive training alone or combined with other interventions (e.g. physical exercise, neuromodulation). We considered cognitive training as an approach that involves a set of standardized tasks designed to maintain or enhance cognitive processes (Simons et al., 2016). Interventions that significantly differ from cognitive training (i.e. not following a series of regular mental activities designed to maintain or improve cognitive performance), such as cognitive behavioural therapy and mindfulness, were excluded.

# Type of outcome measures

The outcome included performance on at least one cognitive test administered both before (baseline) and after the cognitive training program. We consider both standardized instruments and paradigm-based experimental tasks as cognitive tests. Performance improvement was expected in executive functions in neuropsychological tests when comparing baseline pre-training and immediate post-training. In order to employ same-construct comparisons, we categorized the outcome measures by distinct executive functions constituents: inhibitory control, working memory, and cognitive flexibility.

# **Information source**

The following online databases were searched up to April 2021 to identify relevant trials: MEDLINE (PubMed), PsycINFO, The Cochrane Library Cochrane Central Register of Controlled Trials (CENTRAL), Web of Science (Science and Social Science Citation Index), and SciELO (Scientific Electronic Library Online). For the identification and use of descriptors (i.e. specific keywords), we resorted to medical subject headings (MeSH) terms. To include as many trials as possible in addition to the MeSH terms, we included additional descriptors with terms not directly linked to MeSH (called 'Text word'), but closely related to the investigated research topic. Subsequently, a new search was

performed in additional directories up to April 2022 to identify possible updates to previously obtained trials: Epistemonikos (www.epistemonikos.org), Lens (www.lens. org), and Cognitive Training Data (www.cognitivetrainingdata.org). To elaborate the search strategy in the first two directories, we used the 2D Search open-source software (www.2dsearch.com), in which queries are formulated by manipulating objects on a two-dimensional canvas. In the Cognitive Training Data directory, we extracted the trials using Mendeley Reference Manager version 2.63 open-source software (www. mendeley.com). See Supplementary Material—Appendix B for the detailed search strategy for both searches.

# Study selection and risk of bias

There were no restrictions on language and publication date. Two authors (RLO and SRF) independently removed the duplicate items and performed the initial screening (i.e. titles and abstracts reading) of studies identified by the specific search strategy. Divergence in study selection was resolved by the third author (RMJ). The two authors (RLO and SRF) subsequently read the selected studies' full text for potentially eligible studies. We utilized the Rayyan open-source free web-tool software (rayyan.qcri.org) during the entire screening process (on the advantages of Rayyan, see Ouzzani et al., 2016). The two authors (RLO and SRF) collected data about trial identification (title, authors, and year of publication), sample characteristics (sample size, mean age, standard deviation of each group), characteristics of the cognitive training, its duration (sessions), type of control group involved (active or passive), and the outcome measures (inhibitory control, working memory, and cognitive flexibility).

One author (RLO) assessed the methodological quality of the included trials in meta-analysis using the Cochrane Risk of Bias 2 tool (RoB2; for a description, see Sterne et al., 2019). This tool provides a framework for assessing the risk of bias in a single estimate of an intervention effect reported from a trial. RoB2 is structured into seven bias domains (i.e. random sequence generation, allocation concealment, blinding of participants and personnel, blinding of outcome assessment, incomplete outcome data, selective reporting, and other bias).

#### Data extraction and analysis

For the systematic review, the following data were extracted: first author and year, sample, mean age, cognitive training type, dose, sessions, length, sessions per week, executive function outcomes, and control groups. For the meta-analysis, the continuous data to perform the meta-analysis was extracted by one reviewer (RLO) and checked by a second reviewer (RMJ). To perform the meta-analysis, we adopted the Review Manager version 5.4 open-source software (for a description, see Cochrane, 2022). The main outcome was the standardised mean difference (SMD) from pre- to post-training in the experimental group(s) (i.e. cognitive training) and control group(s) (i.e. active or passive).

The analyses on the SMD were conducted for each of the executive function subdomains. Precision of the SMD was calculated for each trial by 95% confidence intervals (CI). The trials were required to have measured participants baseline ability in the trained cognitive skill, and this measure could come from the same task that was later used for training or from a different task that assessed the same cognitive skill. Furthermore, trials were required to have measured participants cognitive training outcomes using gain scores. The continuous data values were entered into a spread-sheet (available at https://osf.io/64xmj/), and then organized to run the meta-analysis.

The selected trials were inserted in a separate spreadsheet tab containing data referring to the pre- and post-intervention, including the sample, mean difference, and standard deviation of the experimental and control groups. Considering that the included trials had distinct populations, intervention parameters, and settings, a random effect model was employed in the meta-analysis. The heterogeneity was assessed by the  $l^2$  statistic and 95% CI. The following  $l^2$  statistics were considered: 0–40%: not important/low heterogeneity; 30%–60%: moderate heterogeneity; 50%–90%: substantial heterogeneity; 75%–100%: considerable heterogeneity (Deeks & Higgins, 2022). The I<sup>2</sup> statistic reports the percentage of variation across selected studies that is due to heterogeneity among studies (from a statistical origin) rather than chance. The higher the  $I^2$ , the higher the likelihood of drawing incorrect metanalytic inferences on the clinical relevance, that is, the effect of the cognitive training (see Melsen et al., 2014). As a complementary measure of heterogeneity, we also calculated the Tau-squared, which measures variance of the true effect sizes. Assessment of clinical relevance was made using three categories: small effect (mean differences [MD] < 10% of the scale; standardised mean difference [SMD] < 0.5); medium effect (MD from 10% to 20% of the scale; SMD from 0.5 to 0.8); large effect (MD > 20% of the scale; SMD > 0.8) (Furlan et al., 2009). In addition, the effect size was assessed by the Cohen's d. Heterogeneity and clinical relevance assessments together allow inferences on the intervention efficacy and enable professionals to decide about the applicability of the results to their target population. A funnel plot for identifying possible publication bias was calculated, and a sensitivity analysis was planned to identify if a specific trial changes the overall effect, by repeating the meta-analysis with one trial omitted at a time (forest plot inspection for outliers). We adopted a significance level of 5% for all tests.

# Results

# Study selection

The initial search in the electronic databases yielded 3,544 trials. After removal of duplicates 2,587 trials were screened. After abstract and title screening, we assessed 75 full-texts for eligibility. We subsequently included 31 trials in the systematic review and 13 trials were selected for meta-analysis. The PRISMA-based flow diagram provides an overview of the trials' selection process (Figure 1).

# Characteristics of the included studies

The characteristics of the individual trials are summarized in Table 1. The publication year of the selected trials ranged from 2009 to 2022, and the participants' age ranged between 59 and 82 years old. The included trials had a total of 2,783 participants of



Figure 1. Flow diagram with data related to trials screening throughout the whole process.

both sexes. The studies extracted from the systematic review had an average (mean ± standard deviation) of  $23.63 \pm 15.69$  total training hours;  $31.10 \pm 19.06$  total training sessions; 3.16±1.42 sessions per week; and 50.63±28.38 min per session. Regarding the type of

	The section of individual mais.										
Authors	Year N	Mean age	CT*	[1]	[2]	[3]	[4]	× ⊇	W C	F AC	PA
Nouchi et al.	2019 EG: 30 CG: 30	EG: 71.67 ± 3.62 CG: 73.11 ± 3.90	T۷	10.00	30	20.00	5 S	> \		>	
Schoene et al.	2015 EG: 47 CG: 43	EG: $82.00 \pm 7.00$ CG: $81.00 \pm 7.00$	۶D	21.95	31	27.40	m	>	>		>
Ballesteros et al.	2014 EG:17 CG:13	EG: 68.80±5.15 CG: 69.20±5.91	9	20.00	20	60.00	2	>	>		>
Shatil et al.	2014 EG: 60 CG: 59	EG: 67.70±5.80 CG: 68.30±5.80	TV	8.00	24	20.00	m	>	>	>	
Reve and Bruin	2014 EG: 76 CG: 69	EG: $81.90 \pm 6.30$ CG: $81.10 \pm 8.30$	9	6.00	36	10.00	m		>	>	
Peretz et al.	2011 EG: 66 CG: 55	EG: 68.6±7.70 CG: 66.9±7.30	9	18.00	36	30.00	m	>	>	>	
Adcock et al.	2020 EG: 15 CG: 16	EG: 77.00 $\pm$ 6.40 CG: 70.90 $\pm$ 5.00	٩G	32.00	48	40.00	ŝ	>	>		>
Simon et al.	2018 EG: 41 CG: 41	EG: $72.40 \pm 5.60$ CG: $73.70 \pm 6.50$	9	16.67	25	40.00	S	>	>	>	
Hardcastle et al.	2022 EG: 30 CG: 28	EG: 70.67 ± 3.99 EC: 71.11 ± 5.28	9	40.00	60	40.00	ŝ	>			>
Nouchi et al.	2012 EG: 14 CG: 14	EG: 68.86±2.07 CG: 69.31±2.82	CP, VG	5.00	20	15.00	S	>	>	>	
Turner et al.	2019 EG: 15 CG: 15	EG: $67.00 \pm 5.87$ CG: $68.08 \pm 4.54$	9	20.00	10	120.00	7	>	>	>	
Grönholm-Nyman et al.	2017 EG: 17 CG: 16	TG: $68.76 \pm 6.68$ CG: $68.31 \pm 8.28$	9	15.00	15	60.00	m	>	>	>	
Kazazi et al.	2021 EG: 26 CG: 26	EG: 65.42 ± 5.40 CG: 64.38 ± 5.00	9	9.00	12	45.00	7	>	>		>
Basak et al.	2008 EG: 19 CG: 20	EG: $70.05 \pm 4.94$ CG: $69.10 \pm 6.06$	۶D	23.5	15	90.00	m	>	>		>
Estrada-Plana et al.	2021 EG: 12 CG: 15	EG: 81.83 ± 8.86 CG: 82.93 ± 8.95	۶	10.00	2	60.00	7	>		>	
Lee et al.	2020 EG: 29 CG:39	EG: 70.41 ± 3.56 CG: 69.69 ± 3.88	₽	35.00	50	42.00	ŝ	>	>	>	
Jaeggi et al.	2020 EG: 78 CG: 77	EG: $72.33 \pm 5.51$ CG: $73.39 \pm 5.33$	9	6.67	20	20.00	7	>			>
Falbo et at.	2016 EG: 20 CG: 16	EG: $71.50 \pm 6.70$ CG: $73.70 \pm 4.50$	Ы	24.00	24	60.00	7	>		>	
Mozolic et al.	2011 EG: 33 CG: 33	EG: $69.40 \pm 3.20$ CG: $69.40 \pm 2.50$	9	8.00	∞	60.00	-	>	>		>
Smith et al.	2009 EG: 242 CG: 245	EG: $75.60 \pm 6.60$ CG: $75.00 \pm 6.30$	8	40.00	40	60.00	S	>		>	
Perrot et al.	2019 EG1:12 EG2:12 CG: 11	EG1: 63.75 ± 2.49 EG2: 64.67 ± 3.17 CG: 65.55 ± 2.91	۶D	24.00	24	60.00	m	>	>		>
Eggenberger et al.	2015 EG1:24 EG2:22 CG: 25	EG1: 77.30±6.30 EG2: 78.50±5.10 CG: 80.80±4.70	CP, VG	34.67	52	40.00	2	>	>	>	
Gajewski et al.	2020 EG: 32 CG1: 33 CG2:37	EG: 71.00 $\pm$ 4.20 CG1: 71.00 $\pm$ 4.50 CG2: 70.00 $\pm$ 4.20	CP, PP	49.07	32	90.00	7		>	>	
Weicker et al.	2018 EG1: 20 CG1:20 CG2:20	EG1: 67.80±3.90 EG2: 67.70±3.10 CG: 67.50±5.70	9	9.00	12	45.00	m	>	>	>	>
Ten Brinke et al.	2019 EG1: 39 EG2: 38 CG: 40	EG: 71.36±5.14 EG2: 72.88±5.17 CG: 72.46±4.11	9	24	24	60.00	m		>	>	
Meltzer et al.	2021 EG1: 28 EG2: 24 CG: 24	EG1: 69.57 ± 2.97 EG2: 70.08 ± 2.89 EG3: 70.00 ± 2.62	SP	40	80	30.00	ŝ	>			>
Nouchi et al.	2021 EG1:36 EG2: 36 CG1: 34 CG2: 36	EG1: 67.97 ± 3.12 EG2: 67.42 ± 4.78 CG1: 67.59 ± 4.58 CG2:	9	21.00	84	15.00	~	>		>	
		$67.86 \pm 4.92$									
Shatil	2013 EG1: 33 EG2: 29 TG3: 29 CG: 29	EG1: 80.00±5.43 EG2: 79.00±5.49 EG3: 81.00±5.25 CG: 70.00+5.76	Ð	32.00	48	40.00	m	>	>	>	
Gaiowski and	2018 EG1·35 EG2·32 EG3·34 CG·40	FG1-71 90+7 40 FG3-70 90+4 10 FG3-71 10+4 50 CG-	dd d)	48.00	37	00.00	<u>ر</u>	>		>	>
Falkenstein		$69.90 \pm 4.20$			1	0	4				
Desjardins-Crépeau	2016 EG1: 22 EG2: 20 CG1: 16 CG2: 18	EG1: $72.70 \pm 7.40$ EG2: $73.20 \pm 6.30$ CG1: $70.90 \pm 7.40$ CG2:	Ð	72.00	36	120.00	m		>	>	
בו מוי		00° / I I C'7 /						`			,
Chen et al.	2017 EG1: 19 EG2: 17 EG3: 15 EG4: 15 CG: 2	0 General: 68.55±5.74	Ы	10.00	10	60.00	-	>			>
Notes. *Type of cogniti of sessions per week	re training (CT) based on the modality of ir Acronyms: Experimental Group (EG), Con	tervention, [1] Total number of training hours, [2] Total number ol Group (CG), TV: TV-Based (stimuli and task delivered on a	of CT sest television	vith no	Single behav	vioural r	length espons	es reg	istered	[4] NL by a	mber v TV
system), VG: Video لاذ	me-Based (stimuli and task adapted to a v	deo game interface with manual motor responses using Joystic	ks and reg	istered b	y a vic	deo gam	e syste	י כ (נו	. Com	puter-	ased

Table 1. The characteristics of individual trials

(stimuli and task in computational interface), PP: Paper-and-pencil-based (analogic stimuli), EC: Ecological Training (tasks in rich and real-world environments related to daily activities), SP: Smartphone App Training (computational stimuli and task adapted to smartphones interface), IC: Inhibitory Control, WM: Working Memory, CF: Cognitive Flexibility (IC, WM, and CF columns indicate whether this mental functions were investigated), AC: Active, PA: Passive.

cognitive training, the selected trials presented approaches based on computer (n=20), videogame (n=7), TV (n=2), paper-and-pencil (n=1), ecological (n=1), and smartphone app (n=1). A total of 21 trials reported an outcome related to inhibitory control, 27 trials related to working memory, and 21 trials related to cognitive flexibility as reported by the columns IC, WM, and CF in Table 1, respectively. Regarding the control groups, 20 trials reported active control group(s) and 13 trials reported passive control group(s)<sup>1</sup>.

The trials were conducted in the United States of America (Basak et al., 2008; Hardcastle et al., 2022; Jaeggi et al., 2020; Lee et al., 2020; Mozolic et al., 2011; Shatil, 2013; Simon et al., 2018; Smith et al., 2009; Turner et al., 2020), Canada (Desjardins et al., 2016; Meltzer et al., 2023; Ten Brinke et al., 2020), Japan (Nouchi et al., 2012, 2019, 2021), Australia (Schoene et al., 2015), Spain (Ballesteros et al., 2014; Estrada-Plana et al., 2021), Switzerland (Adcock et al., 2019; Eggenberger et al., 2015; Van Het Reve & De Bruin, 2014), Israel (Peretz et al., 2011), Germany (Gajewski et al., 2018, 2020; Van Het Reve & De Bruin, 2014; Weicker et al., 2018), France (Perrot et al., 2019), Finland (Grönholm-Nyman et al., 2017), Iran (Kazazi et al., 2021), Italy (Falbo et al., 2016), Sweden (Simon et al., 2018), China (Chiu et al., 2017) and multi-countries (Shatil et al., 2014). The location where the study was carried out (registration in the ethics committee) and location of the first author's affiliation were taken as criteria to establish the study origin.

#### Meta-analyses

Here we present the meta-analyses of each executive function subdomain according to a core triad of functions. Only trials that reported pre- and post-intervention data were included (i.e. MD and standard deviation). Based on the available data for the meta-analysis, we presented the results for the most frequent tasks conducted: Stroop and Go/No-Go tasks for inhibitory control; Digit Span (Forward and Backward), N-Back, and Corsi Blocks tasks for working memory; and Semantic Fluency and Trail Making— Part B tasks for cognitive flexibility.<sup>2</sup>

#### Effects on inhibitory control

The effects of cognitive training on inhibitory control were evaluated in four trials (Kazazi et al., 2021; Nouchi et al., 2019; Perrot et al., 2019; Weicker et al., 2018), which were measured by the Stroop and Go/No-Go tasks. The meta-analysis on inhibitory control measured by the Stroop task resulted in a statistical significance (overall effect) in favour of cognitive training intervention when compared to active/passive controls (Figure 2 upper half; n=143 participants, SMD=.78, CI [.33, 1.22], p < .001, d=1.64). There was low heterogeneity in the overall analysis ( $l^2=35\%$ ,  $Tau^2=.05$ ). Additionally, the meta-analysis resulted in a statistical significance (subgroup effect) in favour of cognitive training when compared to the active control (n=120 participants, SMD=.61, CI [.25, .98], p=.001, d=6.41,  $l^2=0\%$ ,  $Tau^2=.00$ ) and passive control (n=23 participants, SMD=1.53, CI [.58, 2.48], p=.002, d=.41).

The meta-analysis on inhibitory control measured by the Go/No-Go task did not show statistical significance (overall effect) in favour of cognitive training intervention when compared to active/passive controls (Figure 2 bottom half; n=92, SMD=-0.24,

#### Stroop Task

	Expe	erimen	tal	C	ontrol			Std. Mean Difference	Std. Mean Difference	Risk of Bias
Study or Subgroup	Mean	SD	Total	Mean	SD	Stat. Mean Difference     Stat. Mean Difference     Stat. Mean Difference       D Total Weight     IV, Random, 95% CI     IV, Random, 95% CI       1 30 41.1%     0.62 [0.10, 1.14]		ABCDEFG		
1.1.1 Active										
Nouchi et al. (2019)	4	4.88	30	0.89	5.01	30	41.1%	0.62 [0.10, 1.14]		$\bullet \bullet \bullet \bullet \bullet \bullet$
Nouchi et al. (2019)	2.7	3.23	30	0.29	4.48	30	41.2%	0.61 [0.09, 1.13]		$\bullet \bullet \bullet \bullet \bullet \bullet \bullet$
Subtotal (95% CI)			60			60	82.3%	0.61 [0.25, 0.98]	◆	
Heterogeneity: Tau <sup>2</sup> =	0.00; CI	ni² = 0.	00, df=	1 (P =	0.98); (	1 <sup>2</sup> = 0%				
Test for overall effect:	Z = 3.29	(P = 0	.001)							
1.1.2 Passive									1000	
Perrot et al. (2019)	9.25	6.3	12	0.91	3.81	11	17.7%	1.53 [0.58, 2.48]		
Subtotal (95% CI)			12			11	17.7%	1.53 [0.58, 2.48]		
Heterogeneity: Not ap	plicable									
Test for overall effect:	Z = 3.15	(P = 0	.002)							
11012223-02110-01026-022			1000			35-85				
Total (95% CI)			72			71	100.0%	0.78 [0.33, 1.22]	•	
Heterogeneity: Tau <sup>2</sup> =	0.05; CI	ni² = 3.	08, df=	: 2 (P =	0.21);1	I <sup>2</sup> = 359	6			
Test for overall effect:	Z= 3.43	(P = 0	.0006)						-2 -1 U 1 Z	all
Test for subaroup diff	erences	: Chi <sup>2</sup> =	= 3.08.	df = 1 (F	= 0.0	8),   <sup>2</sup> =	37.5%		ravours (control) Favours (experiment	aij

# Go/No-Go Task

	Expe	erimen	tal	C	ontrol			Std. Mean Difference	Std. Mean Difference	Ris	k of Bias
Study or Subgroup	Mean	SD	Total	Mean	SD	Total	Weight	IV, Random, 95% CI	IV, Random, 95% CI	ABC	DEFG
1.2.1 Active											
Weicker et al. (2018) Subtotal (95% Cl)	4.84	6.16	26 26	1.57	5.89	26 26	50.7% 50.7%	0.53 [-0.02, 1.09] 0.53 [-0.02, 1.09]	•		••
Heterogeneity: Not app	licable										
Test for overall effect: Z	= 1.89	(P = 0.1	06)								
1.2.2 Passive									_		
Kazazi et al. (2021)	-0.1	0.66	20	0.4	0.06	20	49.3%	-1.05 [-1.71, -0.38]		• •	••
Subtotal (95% CI)			20			20	49.3%	-1.05 [-1.71, -0.38]			
. Heterogeneity: Not app	licable										
Test for overall effect: Z	= 3.08	(P = 0.1	002)								
Total (95% CI)			46			46	100.0%	-0.24 [-1.79, 1.30]	+		
Heterogeneity: Tau <sup>2</sup> = 1	1.15; Ch	i <sup>2</sup> = 12.	.80, df=	= 1 (P =	0.0003	3); I <sup>2</sup> = 9	2%				
Test for overall effect: Z	= 0.31	(P = 0.1)	76)						-4 -2 U 2 4	an an	
Test for subgroup diffe	rences:	Chi <sup>2</sup> =	12.80,	df = 1 (F	P = 0.0	003), I <sup>z</sup>	= 92.2%		Favours [control] Favours [experimenta	11	
Risk of bias legend											
(A) Random sequence	genera	tion (s	election	n bias)							
(B) Allocation concealm	nent (se	lection	bias)								
(C) Blinding of participa	ants and	perso	nnel (p	erforma	ance b	ias)					
(D) Blinding of outcome	e asses	sment	(detect	tion bia:	s)						
(E) Incomplete outcom	e data (a	attrition	n bias)								
and a 1 11 11 11											

(F) Selective reporting (reporting bias)

(G) Other bias

#### Figure 2. Inhibitory control measured by the Stroop task and Go/No-Go task.

Notes. Standardised mean difference effects of cognitive training compared with active/passive controls on inhibitory control outcomes in healthy older adults measured by the Stroop task (hits in the incongruent condition) and Go/No-Go task (hits in the inhibitory condition). Overall analysis conducted with a random-effects model for the Stroop task (p < .001) and for the Go-No/Go task (p = .76). The diamonds represent pooled standardised mean difference estimate of random-effects meta-analysis; l<sup>2</sup> represents the heterogeneity test; squares represent study-specific estimates; green circles represent low risk of bias; red circles represent high risk of bias; and the empty space represents unclear risk of bias.

Cl [-1.79, 1.30], p = .76, d = .59). There was considerable heterogeneity in the overall analysis of cognitive training on the Go/No-Go task ( $l^2=92\%$ , Tau<sup>2</sup>= 1.15). In an overall analysis (without comparison of subgroups and tests) the meta-analysis did not show statistical significance (n = 118, SMD = .42, Cl [-0.27, 1.11], p = .76,  $l^2 = 84\%$ , Tau<sup>2</sup>= .50).

#### Effects on working memory

The effects of cognitive training on working memory were evaluated in ten trials (Ballesteros et al., 2014; Basak et al., 2008; Grönholm-Nyman et al., 2017; Jaeggi et al., 2020; Kazazi et al., 2021; Lee et al., 2020; Nouchi et al., 2019; Perrot et al., 2019, Shatil et al., 2014; Weicker et al., 2018), which were measured by the Digit Span (Forward and Backward), N-Back, and Corsi Blocks tasks. The meta-analysis on working memory measured by the Digit Span Task—Forward showed a marginal statistical significance (overall effect) in favour of cognitive training intervention when compared to active/ passive controls (see first forest plot in Figure 3; n = 259 participants, SMD = 2.78, Cl [-0.07, 5.62], p = .06, d = .92). There was considerable heterogeneity in the overall analysis ( $l^2 = 98\%$ ,  $Tau^2 = 8.27$ ). The meta-analysis resulted in a statistical significance (subgroup effect) in favour of the cognitive training when compared to the passive control (n = 40 participants, SMD = 2.70, Cl [1.82, 3.57], p < .001, d = .09), although such output was solely based on one trial; and significance was not achieved when cognitive training was compared to active control group (n = 219 participants, SMD = 2.81, Cl [-0.98, 6.61], p = .15, d = .93,  $l^2 = 99\%$ ,  $Tau^2 = 11.10$ ).

The meta-analysis on working memory by Digit Span—Backward did not show significance (overall effect) in favour of cognitive training intervention when compared to active/passive controls (see second forest plot in Figure 3; n = 292 participants, SMD = .45, CI [-2.05, 2.96], p = .72, d = .95). There was also considerable heterogeneity in the overall analysis ( $I^2 = 98\%$ ,  $Tau^2 = 7.99$ ). The meta-analysis resulted in a statistical significance (subgroup effect) in favour of the passive control when compared to the cognitive training (n = 40 participants, SMD = -2.57, CI [-3.43, -1.72], p < .001, d = .06), although such output showed a small size effect and was solely based on one trial.

The meta-analysis on working memory measured by the Corsi Blocks task showed a statistical significance (overall effect) in favour of cognitive training intervention when compared to active/passive controls (see third forest plot in Figure 3; n =133 participants, *SMD*=2.28, *Cl* [.84, 3.73], p = .002, d = .16). There was considerable heterogeneity in the overall analysis ( $l^2 = 90\%$ , Tau<sup>2</sup> = 1.93). This large heterogeneity could be a result of the outcome of the second study of Weicker et al. (2018), since a reduced heterogeneity ( $l^2 = 51\%$ , *Tau<sup>2</sup>* = .18) was observed without the data of Weicker et al. (2018) and the overall effect remained significant. The subgroup effects showed near significance and statistical difference in favour of cognitive training when compared to active (n = 80 participants, *SMD*=3.19, *Cl* [-0.42, 6.79], p = .08, d = .57,  $l^2 = 96\%$ , *Tau<sup>2</sup>*= 6.48) and passive controls (n = 53 participants, *SMD*=1.49, *Cl* [.22, 2.76], p = .02, d = .25,  $l^2 = 75\%$ , *Tau<sup>2</sup>*= .63), respectively.

The meta-analysis on working memory measured by the N-Back task did not show statistical significance (overall effect; see fourth forest plot in Figure 3): n = 442 participants, SMD = -1.62, CI [-3.54, .30], p = .10, d = .26). There was considerable heterogeneity in the overall analysis ( $l^2 = 98\%$ ; Tau<sup>2</sup>= 7.38). Interestingly, results showed a marginal subgroup effect in favour of the active control when compared to the cognitive training (n = 118 participants, SMD = -5.47, CI [-10.98, .04], p = .05, d = 1.04,  $l^2 = 98\%$ ; Tau<sup>2</sup>= 23.02). In an overall analysis (without comparison of subgroups and tests) the meta-analysis did not show statistical significance (n = 565, SMD = .49, CI [-0.60, 1.57], p = .38,  $l^2 = 98\%$ ,  $Tau^2 = 6$ .18).

# Effects on cognitive flexibility

The effects of cognitive training on cognitive flexibility were evaluated in five trials (Grönholm-Nyman et al., 2017; Schoene et al., 2015; Shatil et al., 2014; Simon et al., 2018; Van Het Reve & De Bruin, 2014), which were measured by the Trail Making Task—Part B, and Semantic Fluency tasks. The meta-analysis on cognitive flexibility measured by the Trail Making Task—Part B did not show

#### Digit Span Task - Forward

	Expe	erimen	tal	C	ontrol			Std. Mean Difference	Std. Mean Difference	Risk of Bias
Study or Subgroup	Mean	SD	Total	Mean	SD	Total	Weight	IV, Random, 95% Cl	IV, Random, 95% CI	ABCDEFG
2.1.1 Active										
Nouchi et al. (2019)	0	1.39	30	0.29	1.01	30	25.3%	-0.24 [-0.74, 0.27]		
Shatil et al. (2014)	1.18	0.13	60	0.2	0.14	59	24.7%	7.21 [6.21, 8.21]	-	••
Weicker et al. (2018) Subtotal (95% CI)	0.8	0.2	20 110	0.05	0.65	20 109	25.1% 75.1%	1.53 [0.82, 2.24] 2.81 [-0.98, 6.61]	-	
Heterogeneity: Tau <sup>2</sup> =	11.10; C	hi <sup>2</sup> = 1	59.99,	df = 2 (P	< 0.00	0001);1	= 99%			
Test for overall effect 2	Z= 1.45	(P = 0.1	15)							
2.1.2 Passive										
Weicker et al. (2018) Subtotal (95% Cl)	0.8	0.2	20 20	0.25	0.2	20 20	24.9% 24.9%	2.70 [1.82, 3.57] 2.70 [1.82, 3.57]		
Heterogeneity: Not app	olicable									
Test for overall effect 2	Z = 6.02	(P < 0.1	00001)							
Total (95% CI)			130			129	100.0%	2.78 [-0.07, 5.62]	•	
Heterogeneity: Tau <sup>2</sup> =	8.27; Ch	i² = 173	7.44, di	f=3(P <	< 0.000	001); l²	= 98%			-
Test for overall effect 2	Z = 1.91	(P = 0.0	06)						-20 -10 0 10 20	Intal
Test for subgroup diffe	rences:	Chi2 =	0.00, d	f=1(P:	= 0.95	), $ ^2 = 0$	%		r avours (control) i avours (experime	italj

#### Digit Span Task - Backward

	Expe	rimen	tal	c	ontrol			Std. Mean Difference	Std. Mean Difference	Risk of Bias
Study or Subgroup	Mean	SD	Total	Mean	SD	Total	Weight	IV, Random, 95% CI	IV, Random, 95% CI	ABCDEFG
2.2.1 Active										
Grönholm-Nyman et al. (2017)	0.06	0.13	17	0.31	0.68	16	20.1%	-0.51 [-1.20, 0.19]	-	•• ••
Nouchi et al. (2019)	0.33	0.96	30	0.21	0.99	30	20.2%	0.12 [-0.39, 0.63]	+	
Shatil et al. (2014)	1.12	0.15	60	0.26	0.1	59	19.8%	6.69 [5.76, 7.63]	+	
Weicker et al. (2018) Subtotal (95% CI)	-0.1	0.08	20 127	0.1	0.18	20 125	20.0% 80.1%	-1.41 [-2.11, -0.71] 1.21 [-1.63, 4.05]	*	
Heterogeneity: Tau <sup>2</sup> = 8.25; Chi <sup>2</sup> =	206.67	, df = :	3 (P < 0	.00001	;   <sup>2</sup> = 9	9%				
Test for overall effect: Z = 0.83 (P	= 0.40)									
2.2.2 Passive										
Weicker et al. (2018)	-0.1	0.08	20	0.3	0.2	20	19.9%	-2.57 [-3.43, -1.72]	+	$\bullet \bullet \bullet \bullet \bullet \bullet$
Subtotal (95% CI)			20			20	19.9%	-2.57 [-3.43, -1.72]	♦	
Heterogeneity: Not applicable										
Test for overall effect: Z = 5.88 (P	< 0.000	01)								
Total (95% CI)			147			145	100.0%	0.45 [-2.05, 2.96]	•	
Heterogeneity: Tau <sup>2</sup> = 7.99: Chi <sup>2</sup> =	248.38	3. df = 4	4 (P < 0	.00001	$ ^{2} = 9$	8%			<u></u>	
Test for overall effect: 7 = 0.36 (P)	= 0.72)								-10 -5 0 5 10	
Test for subgroup differences: Ch	ni# = 6.2	4, df =	1 (P = I	0.01), P	= 84.0	%			Favours (control) Favours (experimen	italj

#### Corsi Block Task

	Expe	erimen	ıtal	C	ontrol			Std. Mean Difference	Std. Mean	Difference	F	Risk of Bias
Study or Subgroup	Mean	SD	Total	Mean	SD	Total	Weight	IV, Random, 95% CI	IV, Rando	m, 95% CI	AB	CDEFG
2.4.1 Active												
Weicker et al. (2018)	0.25	0.59	20	-0.35	0.1	20	26.4%	1.39 [0.69, 2.09]		-		
Weicker et al. (2018) Subtotal (95% CI)	1.15	0.36	20 40	-0.25	0.13	20 40	22.8% 49.2%	5.07 [3.75, 6.39] 3.19 [-0.42, 6.79]		•	••	• ••
Heterogeneity: Tau <sup>2</sup> = 6.48	Chi <sup>2</sup> =	23.22.	df = 1 (	P < 0.00	001);1	<sup>2</sup> = 969	6					
Test for overall effect: Z = 1	.73 (P =	0.08)										
2.4.2 Passive												
Ballesteros et al. (2014)	0.04	0.01	17	-23.66	16.4	13	25.2%	2.15 [1.22, 3.08]		+	•	
Perrot et al. (2019) Subtotal (95% CI)	0.67	0.89	12	-0.09	0.83	11	25.6%	0.85 [-0.01, 1.71]		1		
Subtoral (95% CI)			23			24	30.676	1.49 [0.22, 2.70]		•		
Test for overall effect: Z = 2	29 (P =	4.04 0	11 = 1 (P	= 0.04)	r= /:	0%6						
Total (95% CI)			69			64	100.0%	2.28 [0.84, 3.73]		•		
Heterogeneity: Tau <sup>2</sup> = 1.93	; Chi <sup>2</sup> =	29.95,	df = 3 (	P < 0.00	001);1	²= 90%	6		20 10	10 20		
Test for overall effect: Z = 3	.09 (P =	0.002	)						-20 -10 Eavoure (control)	Envolure formarizmenta		
									ravours (control)	r avours jexperintenta		

Test for subgroup differences:  $Chi^2 = 0.002$ Test for subgroup differences:  $Chi^2 = 0.76$ , df = 1 (P = 0.38),  $I^2 = 0\%$ 

N-Back Task										
	Expe	rimenta	al I	C	ontrol			Std. Mean Difference	Std. Mean Difference	Risk of Bias
Study or Subgroup	Mean	SD	Total	Mean	SD	Total	Weight	IV, Random, 95% CI	IV, Random, 95% CI	ABCDEFG
2.3.1 Active										
Grönholm-Nyman et al. (2017)	-76.34	4.93	15	-24.13	7.26	16	11.0%	-8.14 [-10.42, -5.86]	-	
Grönholm-Nyman et al. (20'7)	-86	57.03	15	-47.8	42.44	16	12.8%	-0.74 [-1.48, -0.01]	-	•• ••
Lee et al. (2023)	-177.43	2.82	27	-35.71	25.07	29	12.0%	-7.70 [-9.26, -6.13]	-	
Subtotal (95% CI)			57			61	35.8%	-5.47 [-10.98, 0.04]	-	
Hotorogeneity: Tau <sup>2</sup> = 23.02; Chi	<sup>2</sup> = 87.79,	df = 2 (F	° ≺ 0.00	001); l²	= 98%					
Test for overall effect: Z = 1.94 (F	P = 0.05)									
2.3.2 Passive										
Basak et al. (2308)	-0.1	39.07	19	-49.9	39.15	18	12.8%	1.25 [0.53, 1.96]	-	
Grönholm-Nyman et al. (2017)	9.66	30.38	15	-23.66	16.4	16	12.8%	1.34 [0.55, 2.13]	-	•• ••
Jaeggi et al. (2020)	76	13	76	-43	55	76	12.9%	2.96 [2.50, 3.43]		•• ••
Kazazi et al. (2321)	-89.59	49.03	26	21.99	28.26	26	12.8%	-2.75 [-3.52, -1.97]	•	
Kazazi et al. (2021)	-52.87	37.19	26	-35.63	26.26	26	12.9%	-0.53 [-1.08, 0.03]	1	
Subtotal (95% CI)			162			162	64.2%	0.46 [-1.48, 2.41]	•	
Heterogeneity: Tau <sup>2</sup> = 4.82; Chi <sup>2</sup>	= 189.23,	df = 4 (F	< 0.00	001); l <sup>a</sup> :	= 98%					
Test for overall effect: Z = 0.47 (F	P = 0.64)									
Total (95% CI)			219			223	100.0%	-1.62 [-3.54, 0.30]	•	
Heterogeneity: Tau <sup>2</sup> = 7.38; Chi <sup>2</sup>	= 368.36,	df = 7 (F	< 0.00	001); l <sup>2</sup> :	= 98%				20 10 10 20	
Test for overall effect: Z = 1.65 (F	P = 0.10								Eavoure controll Eavoure formaries	ontall
Test for subgroup differences: C	hi² = 3.95,	df=' (	P = 0.05	5), l <sup>2</sup> = 74	4.7%				, areas journed , areas fashering	o riturij
Fisk of bias legend										
(A) Random sequence generation	on (selecti	on bias	)							
(B) Allocation concealment (sele	ect on bias	)								
(C) Blinding of participants and p	personnel	(perforn	nance t	ias)						
(D) Blinding of outcome assess	ment (dete	ction bi	as)							
White constate as the same date for	And A second second	s								

(E) Incomplete outcome data (attrition k (F) Selective reporting (reporting bias) (6) Other bias

#### Figure 3. Working memory measured by the Digit Span, Corsi Block, and N-back tasks.

Notes. Standardised mean difference effects of cognitive training compared with active/passive controls on working memory outcomes in healthy older adults measured by the Digit Span - Forward and Backward (score), Corsi Block (score), and N-Back (reaction time) tasks. Overall analysis conducted with a random-effects model for the Digit Span Task – Forward (p = .06), Digit Span – Backward (p = .72), Corsi Block (p = .002), and N-Back (p = .10) tasks. The diamonds represent pooled standardised mean difference estimate of random-effects meta-analysis; I<sup>2</sup> represents the heterogeneity test; squares represent study-specific estimates; green circles represent low risk of bias; red circles represent high risk of bias and the empty space represents unclear risk of bias.

statistical significance (overall effect) in favour of cognitive training intervention when compared to active/passive controls (Figure 4 upper half; n = 458 participants, SMD = -.59, CI [-1.25, .08], p = .08, d = .27). There was considerable heterogeneity in the overall analysis ( $l^2 = 91\%$ ;  $Tau^2 = .52$ ). The meta-analysis resulted in a statistical significance (subgroup effect) in favour of the passive control when compared to the cognitive training (n = 81 participants, SMD = -0.53, Cl [-0.98, -0.09], p = .02, d = .33), although such output was solely based on one trial; and significance was not achieved when active control group was compared to cognitive training group (n = 377 participants, SMD = -0.59, CI [-1.44, .26], p =.17, d = .11,  $l^2 = 93\%$ ,  $Tau^2 = .69$ ).

The meta-analysis on cognitive flexibility measured by Semantic Fluency tasks did not show statistical significance (overall effect) in favour of cognitive training intervention when compared to the active controls (Figure 4 bottom half; n = 115participants, SMD = -1.90, CI [-7.28, 3.47], p = .49, d = .06). There was considerable heterogeneity in the overall analysis ( $l^2 = 99\%$ ; Tau<sup>2</sup> = 14.89). In an overall analysis (without comparison of subgroups and tests) the meta-analysis showed a marginal statistical significance (n = 283, SMD = .92, CI [-1.85, .00], p = .05,  $I^2 = 96\%$ ,  $Tau^2 = 1.48$ ).

#### Trail Making Task - Part B

			_							
	Expe	riment	al	C	ontrol			Std. Mean Difference	Std. Mean Difference	Risk of Blas
Study or Subgroup	Mean	SD	Total	Mean	SD	Total	Weight	IV, Random, 95% Cl	IV, Random, 95% Cl	ABCDEFG
3.5.1 Active										
Grönholm-Nyman et al. (2017)	-10.62	7.07	16	-11.4	5.12	15	17.8%	0.12 [-0.58, 0.83]		•• ••
Reve and Bruin (2014)	-25.1	1.9	69	-22.2	2	76	20.9%	-1.48 [-1.85, -1.11]	-	
Shatil et al. (2014)	-18.33	13.15	60	-7.5	1.96	59	20.7%	-1.14 [-1.53, -0.75]	-	••
Simon et al. (2018)	-5.8	4.5	41	-8.5	16.9	41	20.3%	0.22 [-0.22, 0.65]	-	
Subtotal (95% CI)			186			191	79.7%	-0.59 [-1.44, 0.26]		
Heterogeneity: Tau <sup>2</sup> = 0.69; Chi <sup>2</sup> :	= 43.51, (	;) f = 3 (F	° < 0.00	)001); l²	= 93%	5				
Test for overall effect: Z = 1.36 (P	= 0.17)									
3.5.2 Passive										
Schoene et al. (2015)	-3.2	12.3	39	1.4	0.1	42	20.3%	-0.53 (-0.98, -0.09)		
Subtotal (95% CI)			39			42	20.3%	-0.53 [-0.98, -0.09]	•	
Heterogeneity: Not applicable									-	
Test for overall effect: Z = 2.36 (P	= 0.02)									
Total (95% CI)			225			233	100.0%	-0.59 [-1.25, 0.08]	•	
Heterogeneity: Tau <sup>2</sup> = 0.52; Chi <sup>2</sup> :	= 44.62, 0	;lf = 4 (F	° < 0.00	)001); l²	= 91%	6			-4 -2 0 2 4	-
Test for overall effect: Z = 1.73 (P	= 0.08)								Favours (control) Favours (experimen	tall
Test for subgroup differences: C	hi² = 0.01	, df = 1	(P = 0.	90), l² =	0%					,
Semantic Fluen	су Та	isk								
	Exp	erimen	tal	C	ontrol			Std. Mean Difference	Std. Mean Difference	Risk of Bias
Study or Subgroup	Mean	SD	Total	Mean	SD	Total	Weight	IV, Random, 95% Cl	IV, Random, 95% Cl	ABCDEFG
Grönholm-Nyman et al. (2017)	-1.1	0.2	41	1	0.6	41	49.9%	-4.65 [-5.50, -3.80]		••
Simon et al. (2018)	0	0.48	17	-0.93	1.48	16	50.1%	0.84 [0.12, 1.55]	•	
Total (95% CI)			58			57	100.0%	-1.90 [-7.28, 3.47]	<b>•</b>	
Heterogeneity: Tau <sup>2</sup> = 14.89; Cl	ni² = 93.8	8, df = 1	1 (P < 0	.00001)	; <b>I</b> ² = 9	19%			-20 -10 0 10 20	-
Test for overall effect: Z = 0.69 (	(P = 0.49)	)							Favours (control) Favours (experiment	tall
Risk of bias legend										
(A) Random sequence general	tion (sele	ction bi	ias)							

(B) Allocation concealment (selection bias) (C) Blinding of participants and personnel (performance bias)

(D) Blinding of outcome assessment (detection bias) (E) Incomplete outcome data (attrition bias)

(F) Selective reporting (reporting bias)

(G) Other bias

#### Figure 4. Cognitive flexibility measured by the Trail Making Task – Part B and Semantic Fluency tasks.

Notes. Standardised mean difference effects of cognitive training compared with active/passive controls on cognitive flexibility outcomes in healthy older adults measured by the Trail Making Task – Part B (score in seconds) and Semantic Fluency tasks (score in seconds). Overall analysis conducted with a random-effects model for the Trail Making Task – Part B (p = .08) and Semantic Fluency Task (p = .49). The diamonds represent pooled standardised mean difference estimate of random-effects meta-analysis; l<sup>2</sup> represents the heterogeneity test; squares represent study-specific estimates; green circles represent low risk of bias; red circles represent high risk of bias and the empty space represents unclear risk of bias.

# Methodological quality assessment

The methodological quality was assessed in the thirteen trials included in the meta-analysis. The trials of Lee et al. (2020), Nouchi et al. (2019) and Weicker et al. (2018) had the highest score in the seven categories (Figure 5). The categories *incomplete outcome data* and *selective reporting* had a higher percentage of trials with low risk of bias (Figure 6).



Figure 5. Risk of bias summary: Review authors' judgements about each risk of bias item for each included trial.



Figure 6. Risk of bias graph: Review authors' judgements about each risk of bias item presented as percentages across all included studies.

# Discussion

The aim of the present systematic review with meta-analysis was to assess the efficacy of cognitive training on executive functions in healthy older people. The theoretical framework that proposes a core triad of cognitive functions (Diamond, 2013; Lehto et al., 2003; Miyake et al., 2000) was adopted to segment the meta-analysis in its subdomains (inhibitory control, working memory, and cognitive flexibility). We considered different modalities of cognitive training (e.g. multidomain/single domain training and combined interventions). However, we provided a fine-tuned control over the outcomes and, as far as we know, this is the first review with task-specific meta-analytic data targeting executive functions in older people. For the investigation of cognitive training efficacy in inhibitory control, we adopted the Stroop and Go/No-Go tasks; for working memory, we adopted the Digit Span (forward and backward), Corsi Blocks, and N-Back tasks; and for cognitive flexibility, we adopted the Trail Making—Part B and Semantic Fluency tasks. Our search obtained 21 trials that evaluated inhibitory control, 27 trials that evaluated working memory, and 21 trials that evaluated cognitive flexibility in the systematic review. Most of the selected trials also assessed other cognitive functions (e.g. language, processing speed, general cognition). Thirteen trials were considered for the meta-analysis, which resulted in mixed evidence for the subdomains inhibitory control and working memory, and no evidence of improvement for cognitive flexibility.

The evidence regarding the cognitive gains of cognitive training in executive functions is also mixed when we consider the first meta-analyses that investigated the main cognitive functions in healthy older adults. The review conducted by Lampit et al. (2014) selected 29 randomized controlled trials and found no significant effect for executive functions (p = .96; q = .09). Conversely, Chiu et al. (2017) found a small effect size for executive functions (p < .001; q = .42) in 14 randomized controlled trials. Further reviews then broke down executive subdomains in their analyses, and so do we.

#### Effects on inhibitory control

Our task-specific meta-analysis on inhibitory control measured by the Stroop task resulted in a statistical significance (p < .001; d = 1.64) in favour of cognitive training

when compared to controls. Despite the low quantity, the studies considered for analysis on active control showed high methodological quality; and cognitive training improved the outcomes on the Stroop task when compared to both passive and active controls. In addition, low heterogeneity and high effect size were observed. However, cognitive training did not improve the outcomes on the Go/No-Go task. This may be related to the duration of the intervention (see Chiu et al., 2017). The total number of training hours was 52.9% higher in the studies selected for meta-analysis of the Stroop task outcomes compared to the studies selected for the Go/No-Go task outcomes. As new studies are carried out, characteristics of cognitive training such as training duration should be considered in the future meta-analyses.

Two previous reviews segmented the analysis by executive function subdomains (inhibitory control included). Nguyen et al. (2019) examined the efficacy of single- and multidomain cognitive training targeting executive functions in healthy older people. The study was not limited to randomized controlled trials and considered 13 studies for near-transfer inhibitory control outcomes. The results showed a significant difference in favour of the cognitive training and a small effect size (p = .048; g = .13). The study by Wollesen et al. (2020) investigated the effectiveness of cognitive-motor training on executive functions in older people. Eleven randomized controlled trials entered the meta-analysis that showed that the intervention enhanced inhibitory control (*SMD* = .61; p < .001). Taking all the results together, it is highly plausible to assume that inhibitory control can present a mild-to-moderate enhancement by cognitive training in older adults.

# Effects on working memory

We also obtained mixed results on working memory outcomes. The meta-analysis on working memory measured by the Digit Span task—Forward showed a marginal statistical significance in favour of the cognitive training when compared to controls (p = .06, d = .92), and no difference was observed in the Digit Span task—Backwards (p = .72, d = .95), and in the N-Back task (p = .10, d = .26). It is likely that more complex tasks (i.e. Digit Span—Backwards and N-Back) would require more training for any cognitive improvement than simpler tasks (i.e. Digit Span—Forward). Thus, likewise for inhibitory control, the duration of the training may play a role in cognitive gains. The near-transfer observed in the Corsi Blocks task in the present work (p = .002, d =.16) may be seen as an argument for the training duration influence. The total number of training hours was 21.3% higher in the studies selected for meta-analysis of Corsi Blocks task outcomes (significant in favour of the cognitive training) compared to the studies selected for Digit Span—Backward and N-Back task outcomes (non-significant results). In addition, classic designs of Digit Span tasks require free recall of numbers, while classic designs of Corsi Blocks tasks require recognition of spatial locations visually available. Therefore, Corsi Blocks is generally an easier task than Digit Span.

Another route of explanation is related to the nature of the task: verbal vs. spatial. Our results showed an improvement in working memory due to cognitive training in a spatial task (i.e. Corsi Blocks) but not in verbal tasks (i.e. Digit Span—Backward and N-Back<sup>3</sup>). Most influential models of working memory support distinct visuospatial and verbal subsystems (Baddeley, 1986; Luck & Vogel, 1997; Smith et al., 1996) which

are linked to functional disparities. For instance, participants are usually worse in the backward than forward version of the Digit Span, but they have similar performance in backward and forward versions of the Corsi Blocks (Donolato et al., 2017). The distinction between the two subsystems was also observed in its neural basis since visuospatial and verbal working memory activate different brain areas (Chein et al., 2011; Nagel et al., 2013). Considering the literature and our results, it is reasonable to presume that one working memory subsystem may be more prone to enhancement from cognitive training than the other.

Previous reviews also investigated the effect of cognitive training on working memory in healthy older adults. Wollesen et al. (2020) entered five randomized controlled trials in the meta-analysis, and there was no effect on working memory (p = .19). Three other reviews (Chiu et al., 2017; Lampit et al., 2014; Nguyen et al., 2019) showed significant differences favouring the cognitive training over controls but with small effect sizes (20–45 studies entered the meta-analyses; all p < .001 and .22 < g < .35). The meta-analysis of Chiu et al. (2017) considered overall memory outcomes, including short-term memory and working memory; and the meta-analysis of Nguyen et al. (2019) considered short-term memory and working memory. The high heterogeneity observed in the Digit Span and N-Back meta-analysis ( $l^2 = 98\%$ ) hamper the conclusions. Nevertheless, considering our results and previous reviews in older adults, it is plausible to assume that cognitive training causes a near-transfer for some tasks that may present a none-to-mild enhancement in working memory.

#### Effects on cognitive flexibility

Both the meta-analysis on cognitive flexibility measured by the Trail Making task—Part B and Semantic Fluency tasks did not show statistical significance in favour of cognitive training intervention when compared to controls (p > .08; d < .11). Wollesen et al. (2020) entered five randomized controlled trials in the meta-analysis, and also found no effect on cognitive flexibility (p = .09). Nguyen et al. (2019) considered 13 studies, and the analysis revealed a significant difference in favour of the cognitive training and a small effect size for the cognitive flexibility outcomes (p = .031; g = 0.16). The few studies altogether support the idea that cognitive flexibility can present a none-to-mild enhancement by cognitive training in older adults.

# **Clinical considerations**

In the very best scenario considering the current knowledge, cognitive training is likely to provide up to moderate enhancement of executive functions in healthy older adults. Thus, health professionals should whenever possible keep recommending other interventions that have been proven to maintain or boost general cognition and executive functions: physical activity, healthy diet, social and meaningful connections, and daily-life complex cognitive tasks (Dominguez et al., 2021; Kelly et al., 2017; Newson & Kemps, 2006). Nevertheless, despite the modest results, the area must maintain efforts to investigate a low-cost non-invasive and highly flexible behavioural

intervention. Possibly, the plurality of interventions and outcomes seen in cognitive training may produce a fog in the literature of this very recent area.

Although our review summarized studies were carried out in healthy older adults, it may also have implications for clinical populations. For instance, executive dysfunction can be present in the early stages of Parkinson's disease. The most frequent neuropsychological alterations range from mild executive dysfunction in the early stages to mild cognitive impairment, and then dementia in the later stages (Dirnberger & Jahanshahi, 2013). Thus, the implementation of cognitive training focusing on executive functions can be adopted as a complementary treatment in an attempt to delay neuropsychological impairments. Similarly, people living with HIV experience cognitive decline, including in executive functions (Kanmogne et al., 2018). Previous research has shown that cognitive training is associated with improvements in cognitive and daily function among this population, and has suggested that better results occur after longer cognitive training sessions (Wei et al., 2022). Finally, when dealing with older adults, it is necessary to consider the limitations of these individuals when it comes to mobility and sensorial loss, in addition to age-adjusted task complexity and cognitive load.

#### Highlights, limitations, and future directions

This study is the first review with task-specific meta-analytic data targeting executive functions in older people. We believe that a fine-tuned control over the outcomes provides better clarity regarding the efficacy of cognitive training interventions, and hence helps professionals and decision makers who seek for reviews with detailed data and higher degrees of specificity. Given that the world's population is aging faster than ever before and that innovative approaches to improving the quality of life of older people have been investigated in the literature, this study provides an overview of prospective experimental design data (e.g. dose, sessions, length, tests) for the implementation of cognitive training programmes targeting executive functions in healthy older adults, as well as the subdomains of executive functions with greater and less effectiveness, according to the literature. In addition, for the present review we conducted the search in databases and directories not considered in previous reviews with meta-analysis that investigated the effect of cognitive training on executive functions in older adults (Chiu et al., 2017; Lampit et al. 2014; Nguyen et al., 2019; Wollesen et al., 2020); they were: Scielo, Web of Science, Lens.org, epistemonikos. org, and cognitivetrainingdata.org.

Since we had tight control over the cognitive training outcomes in the pre- and post-testing, we ran meta-analyses with few studies. Our study also presents a limitation regarding the types of the selected trials. We have not restricted combined trials (i.e. cognitive training plus a second intervention), and we made no distinction regarding the type of training sessions (simultaneous or sequential training), the duration of the sections, and the cognitive function trained (single-domain or multi-domain training). In addition, the present review only compared data of pre- and post-training. Future investigations may tackle long-term efficacy of task-specific gains in executive function in older people.

Hopefully, as new trials are published and accumulate, it would be optimal to obtain task-specific meta-analytic data segmented by efficacy duration (immediate and long-term), type of transfer (i.e. near and far), and type of outcome. In particular, it would be of great interest to obtain results on ecologically-validated outcomes due to the close relation of executive functions and functional everyday life activities.

# Conclusion

The present review on the effect of cognitive training for healthy older adults showed mixed evidence for inhibitory control and working memory enhancement, and no evidence of improvement for cognitive flexibility.

# Notes

- 1. A passive control group refers to protocol with no intervention between pre- and post-testing (i.e. no-contact control group). An active control group refers to protocol with an intervention, ideally with similar topography of the experimental group's intervention, but planned to promote no enhancement in the cognitive processes of interest that is evaluated in the pre- and post-testing (e.g. playing Sudoku when the experimental intervention seeks to promote gains in inhibitory control) in order to control placebo-like effects of the cognitive training.
- The Stroop task assesses the ability to inhibit cognitive interference, which occurs 2. when the processing of a stimulus feature simultaneously affects the processing of another attribute of the same stimulus (Scarpina & Tagini, 2017). The Go/No-Go task involves a series of decisions in which participants are asked to respond to one class of stimuli, that is, the go stimuli, but not to another class of stimuli, that is, the no-go stimuli (Young et al., 2018). The Digit Span task involves reading out a series of strings of digits to the participants who are required to repeat them in the same or reverse order of presentation (i.e. forward and backward conditions; Tripathi et al., 2019). In the N-Back task participants are presented a series of visual stimuli and they are asked for each stimulus whether it matches a stimulus n positions before, which requires maintaining continuous updating and processing of information (Gajewski et al., 2018). The Corsi Block consists of a surface of scattered blocks in which the examiner taps a sequence of blocks and the participant has to repeat the sequence in the same order or backwards (Kessels et al., 2000). In the Trail Making – Part B, subjects connect 25 encircled numbers and letters in numerical and alphabetical order, alternating between numbers and letters (Linari et al., 2022). In the Semantic Fluency tasks the individuals are required to recall items. Some variations of this test include the fluency of certain classes of words or different semantic categories such as animals and fruits (Lopes et al., 2009).
- 3. From the total of 5 trials that entered the N-Back task meta-analysis, two used pictorial stimuli instead of verbal stimuli (Jaeggi et al., 2020; Kazazi et al., 2021). Both trials compared the cognitive training to passive controls and account for a small treatment effect (SMD = .32).

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#### Data availability statement

Search strategy, data obtained from databases and directories, data extracted for meta-analysis, and outputs of all analyses are available at https://osf.io/64xmj/.

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